

STOCHASTIC STABILITY OF LINEAR SYSTEMS WITH SEMI-MARKOVIAN JUMP PARAMETERS

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Abstract

Over the past two decades considerable effort has been devoted to problems of stochastic stability, stabilisation, filtering and control for linear and nonlinear systems with Markovian jump parameters, and a number of results have been achieved. However, due to the exponential distribution of the Markovian chain, there are many restrictions on existing results for practical applications. In the present paper, we study systems whose jump parameters are semi-Markovian rather than fully Markovian. We consider only linear systems with semi-Markovian jump parameters, and also study systems with phase-type semi-Markovian jump parameters, because the family of phase-type distributions is dense in the families of all probability distributions on $[0, +\infty)$. Some stochastic stability results are obtained. An example is given to show the potential of the proposed techniques.

1. Introduction

In engineering applications, frequently occurring dynamical systems which can be represented by different forms depending on the value of an associated Markov chain process are termed *Markovian jump systems*. Research into this class of systems and their applications has spanned several decades. For some representative work on this general topic, we refer the reader to [3, 7–9, 12–15, 26] and the references therein. Some recent results can be found in [1, 2, 4–6, 10, 16–18, 21–25].

However, because the distribution of the Markov chain is exponential, Markovian jump systems have many limitations/restrictions in applications, and the results obtained for these systems are conservative in some senses. In this paper, our focus is

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on semi-Markovian jump systems. We attempt to achieve similar results for semi-Markovian jump systems to those for Markovian jump systems. To the best of the authors' knowledge, the problems of stability, stabilisation, filtering and control for semi-Markovian jump systems have not yet been investigated.

Due to the density of phase-type distributions in the families of all probability distributions on $[0, +\infty)$, we first define the phase-type semi-Markovian process, then, for simplicity, we study the simplest linear systems with phase-type semi-Markovian jump parameters. Our results are built on the work of [8], but are more meaningful. Some stochastic stability results are obtained. The techniques employed in this paper can be applied to general control and filtering problems including nonlinear systems with semi-Markovian jump parameters.

2. Phase-type semi-Markov processes and Markovisation

Consider a Markov process $r(t)$ on the state space $\{1, 2, \dots, m+1\}$, where the states $1, 2, \dots, m$ are transient and the state $m+1$ is absorbing. The infinitesimal generator is

$$Q = \begin{pmatrix} T & T^0 \\ \mathbf{0} & 0 \end{pmatrix},$$

where the $m \times m$ matrix $T = (T_{ij})$ satisfies $T_{ii} < 0$, $T_{ij} \geq 0$, $i \neq j$, while $T^0 = (T_1^0, T_2^0, \dots, T_m^0)^T$ is a non-negative column vector, and $Te + T^0 = 0$. The initial distribution vector is (\mathbf{a}, a_{m+1}) , where $\mathbf{a} = (a_1, a_2, \dots, a_m)$, $\mathbf{a}e + a_{m+1} = 1$, and e denotes an m -dimensional column vector having 1's as its components.

PROPOSITION 2.1 ([19]). *The distribution of the time at which $r(t)$ is absorbed in $m+1$ is*

$$F(t) = 1 - \mathbf{a} \exp(Tt)e, \quad t \geq 0. \quad (2.1)$$

We recall the following definitions.

DEFINITION 1 ([19]). The state that $r(t)$ reaches at time t is called the phase of the distribution $F(\cdot)$ at time t . The distribution $F(\cdot)$ defined in (2.1) on $[0, +\infty)$ is called a continuous phase-type (PH) distribution and (\mathbf{a}, T) is called its representation of order m .

There are many PH distributions, for example, a negative exponential distribution is a continuous PH distribution, and a k order Erlang distribution E_k is also a continuous PH distribution.

Similar to the study of continuous-time Markov chains, we consider the discrete-time Markov chain on the state space $\{1, 2, \dots, m+1\}$, let the state $m+1$ be absorbing, and denote the transition matrix by

$$P = \begin{pmatrix} T & T^0 \\ \mathbf{0} & 1 \end{pmatrix},$$

where $T = (T_{ij})$ is a quasi-stochastic matrix, $T_{ij} \geq 0$, $Te \leq e$, $T^0 = (I - T)e$ is a column vector, and $I - T$ is non-singular. Define $\mathbf{a} = (a_1, a_2, \dots, a_m) \geq 0$, $a_{m+1} \geq 0$ and $\mathbf{a}e + a_{m+1} = 1$.

DEFINITION 2 ([19]). The discrete distribution taking values on the non-negative integer space P_h is called a discrete PH distribution if and only if it is a distribution of the transition steps when the Markov chain, with transition matrix P and initial distribution (\mathbf{a}, a_{m+1}) , reaches the absorbing state $m+1$. Again (\mathbf{a}, T) is called its representation of order m .

A geometry distribution can be seen as a discrete PH distribution. In general, a continuous PH distribution and a discrete PH distribution are both called PH distributions.

PROPOSITION 2.2 ([19]). *The family of PH distributions is dense in all the families of distributions on $[0, +\infty)$.*

From Proposition 2.2, for every probability distribution on $[0, +\infty)$, we may choose a PH distribution to simulate the original distribution to any accuracy.

REMARK. A PH distribution is the distribution of a hitting time in a finite-state, time-homogeneous Markov chain. In 1954, Jensen [11] first introduced this distribution in an economic model, but no feasible solution was provided. The key that makes a PH distribution a powerful tool is the matrix-analysis method developed by Neuts [20] in 1975. Since the 1960s, PH distributions have been a very effective method for analysing stochastic models in queueing theory, storage theory, reliability theory, etc. They also replace the special status of the negative exponential distribution. But many scientists and engineers are still unfamiliar with PH distributions. Very recently, it was discovered that by virtue of phase-type semi-Markov processes (defined below), the PH distribution has important applications in control theory, Markov decision theory, Markov games and stochastic differential equations. It is believed that more and more applications of the PH distribution will be discovered in many areas in the near future. This paper will show the application of the PH distribution to control theory.

DEFINITION 3. Let E be a finite or countable set. A stochastic process $\hat{r}(t)$ on the state space E is called a phase semi-Markov process or a denumerable phase semi-Markov process (when E is finite, $\hat{r}(t)$ is also called a finite phase semi-Markov process), if the following conditions hold:

- (1) The sample paths of $(\hat{r}(t), t < +\infty)$ are right-continuous step functions and have left-hand limits with probability one.
- (2) Denote the n th jump point of the process $\hat{r}(t)$ by τ_n ($n = 0, 1, 2, \dots$), where $\tau_0 \equiv 0 < \tau_1 < \tau_2 < \dots < \tau_n < \dots, \tau_n \uparrow +\infty$, then τ_n ($n = 0, 1, 2, \dots$) are Markovian with respect to the process $\hat{r}(t)$.
- (3) $F_{ij}(t) = P(\tau_{n+1} - \tau_n \leq t \mid r(\tau_n) = i, r(\tau_{n+1}) = j) = F_i(t)$ ($i, j \in E, t \geq 0$) do not depend on j and n .
- (4) $F_i(t)$ ($i \in E$) is a phase-type distribution.

Obviously, in the case when $F_i(t)$ ($i \in E$) is a negative exponential distribution, the denumerable phase semi-Markov process is a Markov chain. A denumerable phase semi-Markov process is able to overcome the restriction of the negative exponential distribution of the time that a Markov chain spends in any state. However, in much research the key problem of whether a denumerable phase semi-Markov process can replace a Markov chain or not, is equivalent to whether the denumerable phase semi-Markov process can be transformed into a Markov chain or not. And if it is a finite phase semi-Markov process, it can be transformed to a finite Markov chain. The rest of this section will show that the above claim is true.

Let E be a finite or countable nonempty set, $\hat{r}(t)$ be a denumerable phase semi-Markov process on the state space E . Denote the n th jump point of the process $\hat{r}(t)$ by τ_n ($n = 0, 1, 2, \dots$), where $\tau_0 \equiv 0 < \tau_1 < \tau_2 < \dots < \tau_n < \dots$. Let $(a^{(i)}, T^{(i)})$ ($i \in E$) denote the $m^{(i)}$ order representation of $F_i(t)$, where

$$F_i(t) = P(\tau_{n+1} - \tau_n \leq t \mid \hat{r}(\tau_n) = i) \quad (i \in E),$$

$$a^{(i)} = (a_1^{(i)}, a_2^{(i)}, \dots, a_{m^{(i)}}^{(i)}), \quad T^{(i)} = (T_{jk}^{(i)}, j, k \in E).$$

Let

$$p_{ij} = P(\hat{r}(\tau_{n+1}) = j \mid \hat{r}(\tau_n) = i) \quad (i, j \in E),$$

$$P = (p_{ij}, i, j \in E), \tag{2.2}$$

$$(a, T) = \{(a^{(i)}, T^{(i)}), i \in E\}. \tag{2.3}$$

Obviously, the probability distribution of $\hat{r}(t)$ can be determined only by $\{P, (a, T)\}$.

DEFINITION 4. $\{P, (a, T)\}$ is called the pair of a denumerable phase semi-Markov process $\hat{r}(t)$. For every n ($n = 0, 1, \dots$), $\tau_n \leq t < \tau_{n+1}$, define $J(t) =$ the phase of $F_{\hat{r}(t)}(\cdot)$ at time $t - \tau_n$.

DEFINITION 5. The term $J(t)$ defined above is called the phase of $\hat{r}(t)$ at time t . For any $i \in E$, we define

$$T_j^{(i,0)} = 1 - \sum_{k=1}^{m^{(i)}} T_{jk}^{(i)} \quad (j = 1, 2, \dots, m^{(i)}),$$

$$G = \{(i, k^{(i)}) \mid i \in E, k^{(i)} = 1, 2, \dots, m^{(i)}\}.$$

From the above analysis, we can easily get the following result.

THEOREM 2.3. $Z(t) = (\hat{r}(t), J(t))$ is a Markov chain with state space G (G is finite if and only if E is finite). The infinitesimal generator of $Z(t)$ given by $Q = (q_{\mu\nu}, \mu, \nu \in G)$ is determined only by the pair of $\hat{r}(t)$ given by $\{P, (\alpha, T)\}$ as follows:

$$\begin{cases} q_{(i,k^{(i)})(i,k^{(i)})} = T_{k^{(i)}k^{(i)}}^{(i)}, & (i, k^{(i)}) \in G, \\ q_{(i,k^{(i)})(i,\bar{k}^{(i)})} = T_{k^{(i)}\bar{k}^{(i)}}^{(i)}, & k^{(i)} \neq \bar{k}^{(i)}, (i, k^{(i)}) \in G \text{ and } (i, \bar{k}^{(i)}) \in G, \\ q_{(i,k^{(i)})(j,k^{(j)})} = p_{ij} T_{k^{(i)}k^{(j)}}^{(i,0)} a_{k^{(j)}}^{(j)}, & i \neq j, (i, k^{(i)}) \in G \text{ and } (j, k^{(j)}) \in G. \end{cases} \quad (2.4)$$

Assume that G has $s = \sum_{i \in E} m^{(i)}$ elements, so the state space of $Z(t)$ has s elements. We number the s elements according to the following method: denote the number of (i, k) by $\sum_{r=1}^{i-1} m^{(r)} + k$ ($1 \leq k \leq m^{(i)}$). Also denote this transformation by φ . Hence one has

$$\varphi((i, k)) = \sum_{r=1}^{i-1} m^{(r)} + k \quad (i \in E, 1 \leq k \leq m^{(i)}). \quad (2.5)$$

Moreover, we define

$$\begin{aligned} \alpha_{\varphi((i,k))\varphi((i',k'))} &= q_{(i,k)(i',k')}, \\ r(t) &= \varphi(Z(t)). \end{aligned} \quad (2.6)$$

Therefore $r(t)$ is a Markov chain with state space $S = \{1, 2, \dots, s\}$ and infinitesimal generator $Q = (\alpha_{im}, 1 \leq i, m \leq s)$. We end this section by giving the following definition.

DEFINITION 6. The Markov chain $r(t)$ is called the associated Markov chain of $\hat{r}(t)$.

3. Problem statement and main results

Let $\{\hat{r}(t), t \geq 0\}$ be a finite phase semi-Markovian chain with state space E and probability distribution $\hat{r}(t)$ which can be determined by $\{P, (\alpha, T)\}$ as defined in

(2.2) and (2.3). Consider a class of stochastic linear systems with semi-Markovian jump parameters in a fixed probability space (Ω, \mathcal{F}, P) :

$$\begin{cases} \dot{x}(t) = \hat{A}(\hat{r}(t))x(t), & t \geq 0, \\ x(0) = x_0, \end{cases} \tag{3.1}$$

where the initial state x_0 is a fixed nonrandom constant vector, $r(0) = r_0$ and $\hat{A}(i)$ ($i \in E$) are some known matrices.

For any $i \in E, k = 1, 2, \dots, m^{(i)}$, we define matrices A as follows:

$$A(\varphi(i, k)) \triangleq \hat{A}(i),$$

where φ is defined in (2.5).

It is easy to show that for any $(\omega, t), A(r(\omega, t)) \equiv \hat{A}(\hat{r}(\omega, t))$. Subsequently, we have the following theorem.

THEOREM 3.1. *System (3.1) is equivalent to the following system:*

$$\begin{cases} \dot{x}(t) = A(r(t))x(t), & t \geq 0, \\ x(0) = x_0, \end{cases} \tag{3.2}$$

where $r(t)$ is the associated Markov chain of $\hat{r}(t)$.

System (3.2) has been extensively studied in [8]. By use of Theorem 3.1 and the results of [8], we have the following main result of this paper.

THEOREM 3.2. *For system (3.1),*

- (a) *the asymptotic mean square stabilities, the exponential mean square stability and the stochastic stability are equivalent;*
- (b) *all second moment stability imply almost sure stability;*
- (c) *a necessary and sufficient condition for stochastic stability is that there exist positive definite matrices M_j for $j = 1, 2, \dots, s$ such that*

$$-\alpha_i M_i + \sum_{j \neq i} \alpha_{ij} M_j + A_i' M_i + M_i A_i = -I, \tag{3.3}$$

where α_{ij} is defined in (2.6), $\alpha_i = -\alpha_{ii} = \sum_{j \neq i} \alpha_{ij}$, $A_i = A(i)$ and A_i' is the transpose of A_i , $i, j = 1, 2, \dots, s$.

REMARK. It is worth mentioning that the advantages of Theorem 3.2 are that when studying stochastic stability problems, we can replace Markovian jump systems with semi-Markovian jump systems, and get the same results, while semi-Markovian jump

systems operate under fewer restrictions and can be widely found and used in many real system applications. Also, more importantly, almost all the nice results obtained so far for Markovian jump systems, for example, Ji and Chizeck ([12, 13]), Ji, *et al.* ([14]), Shi and Boukas ([21]), Shi, *et al.* ([22–24]), de Souza and Fragoso ([25]), and the references therein, are valid for semi-Markovian jump systems.

4. Example

We consider the one-dimensional case, that is, $x(t)$ takes values in $R = (-\infty, +\infty)$. Furthermore, we assume that the phase semi-Markovian process $\hat{r}(t)$ has two states denoted by 1 and 2. The sojourn time in the first state is a random variable with negative exponential distribution with parameter λ_1 . The sojourn in the second state is divided into two parts; the sojourn times in the two parts are two random variables which are negative exponentially distributed with parameters λ_2 and λ_3 , respectively. In particular, if the process $\hat{r}(t)$ enters state 2, it must first stay in the first part of state 2 for some time and then stay in the second part, before finally entering state 1. We assume that $\hat{A}_1 \neq \hat{A}_2$ and $p_{12} = p_{21} = 1$.

Obviously,

$$P = \begin{pmatrix} p_{11} & p_{12} \\ p_{21} & p_{22} \end{pmatrix} = \begin{pmatrix} 0 & 1 \\ 1 & 0 \end{pmatrix}, \quad T^{(2)} = \begin{pmatrix} T_{11}^{(2)} & T_{12}^{(2)} \\ T_{21}^{(2)} & T_{22}^{(2)} \end{pmatrix} = \begin{pmatrix} -\lambda_2 & \lambda_2 \\ 0 & -\lambda_3 \end{pmatrix},$$

$$a^{(1)} = (a_1^{(1)}) = (1), \quad T^{(1)} = (T_{11}^{(1)}) = (-\lambda_1), \quad a^{(2)} = (a_1^{(2)}, a_2^{(2)}) = (1, 0).$$

The state space of $Z(t) = (\hat{r}(t), J(t))$ is clearly $G = ((1, 1), (2, 1), (2, 2))$. We number all the elements of G as follows:

$$\varphi((1, 1)) = 1, \quad \varphi((2, 1)) = 2, \quad \varphi((2, 2)) = 3.$$

Hence the infinitesimal generator of $\varphi(Z(t))$ is

$$Q = (\alpha_{ij}) = \begin{pmatrix} -\lambda_1 & \lambda_1 & 0 \\ 0 & -\lambda_2 & \lambda_2 \\ \lambda_3 & 0 & -\lambda_3 \end{pmatrix}.$$

As $x(t) \in R$, so A_i, M_i ($i = 1, 2, 3$) are real constant numbers, and $A'_i = A_i, A_1 = \hat{A}_1, A_2 = A_3 = \hat{A}_2$. Hence (3.3) reduces to

$$\begin{aligned} -\lambda_1 M_1 + \lambda_1 M_2 + 2M_1 \hat{A}_1 &= -1, \\ -\lambda_2 M_2 + \lambda_2 M_3 + 2M_2 \hat{A}_2 &= -1, \\ \lambda_3 M_1 - \lambda_3 M_3 + 2M_3 \hat{A}_2 &= -1. \end{aligned} \tag{4.1}$$

Define $\Delta = (\lambda_1 - 2\hat{A}_1)(\lambda_2 - 2\hat{A}_2)(\lambda_3 - 2\hat{A}_2) - \lambda_1\lambda_2\lambda_3$. From (4.1), if $\Delta \neq 0$, then

$$M_1 = \frac{(\lambda_2 - 2\hat{A}_2)(\lambda_3 - 2\hat{A}_2) + \lambda_1\lambda_2 + \lambda_1(\lambda_3 - 2\hat{A}_2)}{\Delta}, \tag{4.2}$$

$$M_2 = \frac{(\lambda_1 - 2\hat{A}_1)(\lambda_3 - 2\hat{A}_2) + \lambda_2\lambda_3 + \lambda_2(\lambda_1 - 2\hat{A}_1)}{\Delta}, \tag{4.3}$$

$$M_3 = \frac{(\lambda_1 - 2\hat{A}_1)(\lambda_2 - 2\hat{A}_2) + \lambda_1\lambda_3 + \lambda_3(\lambda_2 - 2\hat{A}_2)}{\Delta}. \tag{4.4}$$

If $\Delta = 0$, then $\lambda_1 - 2\hat{A}_1 \neq 0, \lambda_2 - 2\hat{A}_2 \neq 0, \lambda_3 - 2\hat{A}_2 \neq 0$. Hence from (4.1) we have

$$M_1 = \frac{1 + \lambda_1 M_2}{(\lambda_1 - 2\hat{A}_1)}, \quad M_2 = \frac{1 + \lambda_2 M_3}{(\lambda_2 - 2\hat{A}_2)}, \quad M_3 = \frac{1 + \lambda_3 M_1}{(\lambda_3 - 2\hat{A}_2)}. \tag{4.5}$$

Therefore we have the following result.

(1) If $\hat{A}_1 < 0, \hat{A}_2 < 0, \hat{A}_1 \neq \hat{A}_2$, then $M_1 > 0, M_2 > 0, M_3 > 0$. Then, by Theorem 3.1, system (3.1) is stable. In other words, if the original two systems are stable, then the new system, which is generated by the original two systems switched by a phase-type semi-Markovian process, is still stable.

(2) If $\hat{A}_1 > 0, \hat{A}_2 > 0, \lambda_1 > 2\hat{A}_1, \lambda_2 > 2\hat{A}_2, \hat{A}_1 \neq \hat{A}_2$, then $M_1 < 0, M_2 < 0, M_3 < 0$. By Theorem 3.2, system (3.1) is not stable. In other words, if the original two systems are unstable, the new system, which is generated by the original two systems switched by a phase-type semi-Markovian process, is likely to be unstable as well.

(3) If $\hat{A}_1 = -10, \hat{A}_2 = 3, \lambda_1 = 10, \lambda_2 = 20$, then $M_1 = 0.2851, M_2 = 0.7553, M_3 = 0.4787$. By Theorem 3.2, system (3.1) is stable. On the other hand, if $\hat{A}_1 = -10, \hat{A}_2 = 6, \lambda_1 = 10, \lambda_2 = 20$, then $M_1 = -0.1654, M_2 = -0.5962, M_3 = -0.2885$. By Theorem 3.2, the new system (3.1) is unstable. Hence, if one of the two original systems is stable and the other is unstable, then the new switching system, which is generated by the original two systems switched by a phase-type semi-Markovian process, is stable or unstable depending on the parameters.

For the general situation, by Theorem 3.2 and (4.2)–(4.5), we present our last result in this paper.

COROLLARY 4.1. *Consider system (3.1) under the conditions given above.*

(i) *Assume that $\Delta \neq 0$, then system (3.1) is stable if and only if the following three inequalities hold:*

$$\begin{aligned} \Delta \times [(\lambda_2 - 2\hat{A}_2)(\lambda_3 - 2\hat{A}_2) + \lambda_1\lambda_2 + \lambda_1(\lambda_3 - 2\hat{A}_2)] &> 0, \\ \Delta \times [(\lambda_1 - 2\hat{A}_1)(\lambda_3 - 2\hat{A}_2) + \lambda_2\lambda_3 + \lambda_2(\lambda_1 - 2\hat{A}_1)] &> 0, \\ \Delta \times [(\lambda_1 - 2\hat{A}_1)(\lambda_2 - 2\hat{A}_2) + \lambda_1\lambda_3 + \lambda_3(\lambda_2 - 2\hat{A}_2)] &> 0. \end{aligned}$$

(ii) Assume that $\Delta = 0$, then system (3.1) is stable if and only if the following three inequalities hold:

$$(\lambda_1 - 2\hat{A}_1) > 0, \quad (\lambda_2 - 2\hat{A}_2) > 0, \quad (\lambda_3 - 2\hat{A}_2) > 0.$$

5. Conclusions

In this paper, a new method has been proposed to study the problem of stochastic stability for systems with semi-Markovian jump parameters. It is a first attempt on this topic. It is shown that the existing results for stochastic stability for Markovian jump systems also hold for semi-Markovian jump systems. However semi-Markovian jump systems are less conservative and more applicable. A numerical example is given to illustrate the feasibility and effectiveness of the theoretic results obtained. The approach developed in this paper can also be applied to general nonlinear systems with semi-Markovian jump parameters.

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